Investor Beliefs and State Price Densities in the Crude Oil Market*

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Abstract

Standard asset pricing theory suggests that state price densities (SPDs or the pricing kernel) monotonically decrease with returns. We find that the SPDs implicit in the crude oil market display a time varying U-shape pattern. This implies that investors assign high state prices to both negative and positive returns. We use data of the crude oil derivatives market, where speculation and short sales are not regulated, to document how the SPDs are dependent on investor beliefs. Investors' preferences over return states are reflected in belief dispersions in options and underlying futures contracts. Investors overall assign higher state prices to negative returns when there are higher demands for out-of-the-money put options, and when there are increased speculations in the crude oil market after 2004-2005.

JEL Classification: G12, G13

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1 Introduction

State price densities (SPDs or the pricing kernel) contain information about investor preferences and essentially determine expected returns and risk premia. According to the no-arbitrage principle, we can price any asset as long as we know the SPDs and the final payoff of this asset. Although several studies have estimated the SPDs using equity market data,[1] few papers investigate other financial markets.[2] Given that (a) commodities have emerged as a fast growing asset class, (b) each market only contains part of the wealth of the aggregate economy and a subset of information about the aggregate pricing kernel, and (c) participants in the commodity market are often different from investors in other markets, a systematic study of the SPDs implicit in this market is warranted. Using the crude oil market, which is the largest and most liquid commodity derivatives market, this paper estimates the SPDs and backs out investor preferences towards different states of nature.

We estimate SPDs using crude oil futures and options data from 1990 to 2012, and we document time variation in the SPDs and their dynamic structure. We find that the SPDs in the crude oil market display a time varying asymmetric U-shape pattern: Investors assign high state prices to both negative and positive returns. The slope of the U-shaped SPDs in both the decreasing and increasing regions also varies depending on the market condition. In other words, investors' marginal value of an one-dollar payoff in certain states, when returns are either negative or positive, exhibits significant changes across time. We also document that average returns on out-of-the-money (OTM) call and put options are negative, consistent with the SPDs having both decreasing and increasing regions.

Another strand of the literature has argued that SPDs depend on differences in investor beliefs, and that this heterogeneity affects expected returns and the price of risk (e.g., Anderson, Ghysels, and Juergens, 2005; Beber, Buraschi, and Breedon, 2010). In particular, Bakshi, Madan, and Panayotov (2010) advocate that the pattern of the SPDs implied by the index option is compatible with theory when investors have heterogeneous beliefs and are able to short sell equities. However, there is limited literature directly testing how heterogeneous beliefs affect the shape of the SPDs due to various reasons: Heterogeneity of investor beliefs is difficult to measure precisely; short selling is highly regulated in some financial markets, while it is a critical element for the theory to explain the empirical SPDs; and accurate estimation of the dynamic structure of the SPDs can be data-demanding.

[1] An incomplete list includes Aït-Sahalia and Lo (2000), Jackwerth (2000), Rosenberg and Engle (2002), Chernov (2003), Chabi-Yo (2012), Christoffersen, Heston, and Jacobs (2013), Linn, Shive and Shumway (2014), Song and Xiu (2014).

[2] Notable exceptions are Beber and Brandt (2006) and Li and Zhao (2009) who study empirical SPDs in the fixed income market.

The crude oil market provides an excellent laboratory to examine the dependence of SPDs on investor beliefs for the following reasons. First, data on speculators' positions are available in this market. The level of speculation can be interpreted as a measure of heterogeneous beliefs because speculators usually bet on certain price movements, and disagreements among investors induce speculative trading (e.g., Scheinkman and Xiong, 2003). Second, there are no restrictions on short sales in this market. While some investors trade derivatives to hedge risks according to their real demand and supply of crude oil, others may take any positions simply based on their beliefs. Therefore we can test whether investor beliefs affect the SPDs. Third, the crude oil derivatives market is fast growing, and historical data are available for more than twenty years. We have large cross-sections of derivative prices, which allow us to accurately extract the SPDs. Moreover, data on trading volumes and open interests of both futures and options enable us to construct various measures of investor heterogeneous beliefs as suggested by the literature (e.g. Kandel and Pearson, 2005; Buraschi and Jiltsov, 2006).

We investigate whether investor beliefs embedded in trading activities in crude oil futures and options affect the slope of the SPDs, one of the fundamental characteristics of the SPDs. The slope can be related to investors' risk aversion (Rosenberg and Engle, 2002). It also compares the marginal value of an one-dollar payoff in different economic states, as measured by the level of futures returns. Since the physical densities of oil futures returns are relatively symmetric, the slope of the SPDs is primarily characterized by risk-neutral skewness. If risk-neutral skewness is more negative (positive), the decreasing (increasing) region of the SPDs has a steeper slope, and investors assign higher state prices to negative (positive) returns. We first provide evidence that a direct measure of the slope of the SPDs is affected by investor beliefs reflected in options and futures contracts. The slope itself, however, can be noisy around the distribution tails. We therefore focus on risk-neutral skewness, which is a more reliable and comprehensive measure of variations in the marginal rate of substitution across states. We calculate risk-neutral skewness from two distinct approaches and regress it on investor beliefs. To highlight how the SPDs depend on market participation of both hedgers and speculators, we run regressions in two sub-periods: from 1990 to 2004, a period marked by relatively stable participation and prices, and from 2005 to 2012, a period marked by extreme price movements and increased participation by various types of investors, especially speculators. The selection of these sub-sample periods is based on the evidence by Tang and Xiong (2012) and Singleton (2014), who document the financialization of commodity markets around 2004-2005 and how speculative investments affect commodity futures prices.[3]

Our empirical results indicate that the slope of the SPDs is affected by various measures

[3] More recent contributions on the financialization of commodity markets include Cheng and Xiong (2014), Henderson, Pearson and Wang (2014) and others.

of the heterogeneity of investor beliefs. When there are more OTM put option demands, the risk-neutral distribution is more negatively skewed. As such, the decreasing side of the SPD has a steeper slope and investors assign higher state prices to negative returns. Regressions for two sub-periods reveal that more intensive speculation after 2005 is associated with more negative risk-neutral skewness. This is consistent with Singleton (2014) who finds that large fund flows may cause high crude oil futures prices, which implies a high probability of negative expected returns.

We also investigate the impact of investor beliefs of the equity market and the aggregate economy on the SPDs. Although investors' expectations about the equity market significantly affect index option prices and return distributions in the equity market (Han, 2008), bearish stock market beliefs do not imply more negative risk-neutral skewness of crude oil futures returns, and do not have significant impacts on the slope of the SPDs in the crude oil market. It is the participants in the crude oil market, rather than the investors in other markets, who assign marginal values to certain states; only oil market-specific belief measures significantly affect state prices of returns in the crude oil market. Our findings align with Goldstein, Li and Yang (2013), who argue that although information is integrated and fast moving, financial markets can be relatively segmented due to the specialization and friction of investments.

This paper is part of a growing list of recent studies that examine how the activities of investors in the commodity market, both hedgers and speculators, affect futures prices and returns. Hamilton and Wu (2014) document significant changes in oil futures risk premia due to active investments from financial institutions in recent years. Motivated by the coincident price rise and increased financial participation in the crude oil market, Büyükşahin and Harris (2011) analyze whether the crude oil price is driven by hedge funds and other speculators. Acharya, Lochstoer, and Ramadorai (2013) find that producers' hedging demand, captured by their default risk, predicts commodity returns. Hong and Yogo (2012) show that the high level of commodity market activity, measured by the high open-interest growth, predicts high commodity returns. Etula (2013) finds that the supply of speculator capital, captured by changes in broker-dealer balance sheets, predicts commodity returns, especially in energy commodities. However, this paper examines how investor beliefs embedded in trading activities affect the SPDs, which determine not only the commodity futures prices and returns, but also the risk premia.

The rest of the paper proceeds as follows. We present the estimation of the risk-neutral densities, physical densities, and the SPDs in Section 2. Section 3 discusses how the SPDs are affected by investor beliefs. Section 4 concludes.

2 Risk-Neutral Densities and SPDs in the Crude Oil Market

In this section, we discuss the economic framework to obtain the SPDs in the crude oil market based on the no-arbitrage principle. We then discuss the estimation methodology of the risk-neutral densities, physical densities and the SPDs using futures and option prices. Next we describe futures and option data and we present the estimated SPDs, as well as option implied moments.

2.1 Economic Framework

If we denote the SPD by ξ, based on the no-arbitrage principle as long as we know the final payoff p_T of an asset, the price of the asset at time t can be obtained by

$$p_t = E[\xi_T p_T | \mathcal{F}_t], \tag{1}$$

where $\mathcal{F}_t$ denotes the investors' information set at time t. Suppose at time t we have a European call option written on a futures contract $F_{t,T}$ with the strike price K, which matures at time T.[4] Its price is the final payoff discounted to time t,

$$\begin{aligned} C(F_{t,T}, K, t, T) &= E[\xi_T (F_T - K)^+ | \mathcal{F}_t] \\ &= \int_K^\infty \xi_T(x)(F_T(x) - K) P(F_T(x)|\mathcal{F}_t) dx, \end{aligned} \tag{2}$$

where we use $P(F_T(x)|\mathcal{F}_t)$ to denote the conditional physical density at time t. However, note that the general SPD or pricing kernel depends on many state variables and is unknown to investors.[5]

In order to price derivatives, we usually rely on the price dynamics of underlying assets under the risk-neutral measure Q. Under this measure, option prices discounted at the riskless rate are martingales. At time t, we price the call option by

$$\begin{aligned} C(F_{t,T}, K, t, T) &= e^{-r(T-t)} E^Q[(F_T - K)^+ | \mathcal{F}_t] \\ &= e^{-r(T-t)} \int_K^\infty (F_T(x) - K) P^Q(F_T(x)|\mathcal{F}_t) dx, \end{aligned} \tag{3}$$

[4] Crude oil options expire three business days prior to the expiration of the underlying futures contract. To simplify the notation, we do not explicitly distinguish between the futures maturity date T and the option maturity date T' in this paper.

[5] Since we estimate the SPDs using certain specific assets, which are only a subset of the aggregate wealth, we can only obtain the SPDs projected onto those assets. For example, the SPDs estimated from index options are the projection of ξ onto the index returns.

where $P^Q(F_T(x)|\mathcal{F}_t)$ is the conditional density of F_T under the risk-neutral measure. Based on this equation, we can price any option with a known final payoff once we have $P^Q(F_T(x)|\mathcal{F}_t)$. Breeden and Litzenberger (1978) have shown that $P^Q(F_T(x)|\mathcal{F}_t)$ can be obtained by taking the second order derivative of call prices with respect to the strike price K,

$$P^Q(F_T|\mathcal{F}_t) = e^{r(T-t)}\frac{\partial^2 C(F_{t,T}, K, t, T)}{\partial K^2}|_{K=F_T}. \tag{4}$$

Although it is not possible to obtain the SPD ξ that is defined over the aggregate economy, we estimate the SPD in the crude oil market, $\widetilde{\xi}$, and we focus on this specific but relatively segmented financial market. This allows us to infer relevant information about investors' preferences and expectations for the purpose of pricing crude oil derivatives; and how investors value certain economic states and their expectations about the probability of those states in the crude oil market. Combining equations (2) and (3), we get the SPD in the crude oil market,

$$\widetilde{\xi}(F_T|\mathcal{F}_t) = e^{-r(T-t)}\frac{P^Q(F_T|\mathcal{F}_t)}{P(F_T|\mathcal{F}_t)}. \tag{5}$$

Defined as the Arrow-Debreu price of per unit of probability, SPDs reflect how investors evaluate possible states of nature and their expectations of the probability of those states happening. While many studies estimate the SPDs using index options (i.e., $\widetilde{\xi}$ in the equity market [e.g., Jackwerth, 2000]) and interest rate derivatives (i.e., $\widetilde{\xi}$ in the fixed income market [e.g., Li and Zhao, 2009]), this paper investigates the SPDs in the crude oil market. As shown in (5), SPDs depend on two components: risk-neutral densities and physical densities. We first discuss estimation of the risk-neutral density.

2.2 Estimation of the Risk-Neutral Density

We compute conditional estimates of the risk-neutral density using option prices on a particular day. More specifically, we adapt the semi-parametric approach introduced by Aït-Sahalia and Lo (1998) and further developed by Christoffersen and Jacobs (2004). The semi-parametric approach is designed to utilize all available information implicit in the entire cross-section of option prices, while keeping the parametric assumptions to a minimum. On a given day, we first fit Black (1976) implied volatilities of the cross-sectional option data as a second order polynomial function of strike price and maturity. Then we construct a grid of strike prices and obtain at-the-money Black (1976) implied volatilities from the fitted polynomial function for each maturity $T-t$. With these implied volatilities, we back out call prices $\widehat{C}\left(F_t, K, t, T, \sigma(K,T)\right)$ on the desired grid of strike prices, and then calculate the risk-neutral density (4) for the spot price at the maturity date T. Lastly we compute the second order derivative of the fitted option price with respect to the strike price

$$\widehat{P}^{Q}(F_T|\mathcal{F}_t) = e^{r(T-t)} \cdot \frac{\partial^2 \widehat{C}\left(F_{t,T}, K, t, T, \sigma(K,T)\right)}{\partial K^2}|_{K=F_T}. \tag{6}$$

Let the return of longing a futures contract maturing at T be $R_{t,T} = \log(F_T/F_{t,T})$. It is the same as $\log(S_T/F_{t,T})$, since the futures price eventually converges to the spot price. We can obtain the density of futures return over the period of $T-t$ as

$$\begin{aligned}\widehat{P}^{Q}(R_{t,T}|\mathcal{F}_t) &= \frac{\partial}{\partial u}\Pr\left(\log(F_T/F_{t,T}) \le u|\mathcal{F}_t\right) = \frac{\partial}{\partial u}\Pr\left(F_T \le F_{t,T}\exp\left(u\right)|\mathcal{F}_t\right) \\ &= \widehat{P}^{Q}(F_{t,T}\exp\left(u\right)|\mathcal{F}_t) \cdot F_{t,T}\exp\left(u\right), \end{aligned} \tag{7}$$

where Pr(.) denotes the cumulative distribution function. We compute a fixed one-month (21 business days) horizon option-implied density by interpolating the term structure of density (7) on each day. Alternatively, we obtain the risk-neutral density of futures returns using the nonparametric approach from Aït-Sahalia and Duarte (2003) and Li and Zhao (2009), which we have not reported here. Results are qualitatively similar. Our estimation spans a long period while keeping the parametric assumptions to a minimum, compared with the existing literature in the commodity market (e.g., Melick and Thomas [1997] estimate the risk-neutral distribution from crude oil options around the time of the first Gulf war under restrictive lognormal assumptions).

2.3 Estimation of the Physical Density and the SPD

Once we have obtained the risk-neutral density $\widehat{P}^{Q}(R_{t,T}|\mathcal{F}_t)$ from option prices, the other component needed to compute the SPD is the physical density $\widehat{P}(R_{t,T}|\mathcal{F}_t)$. We estimate the physical density from the historical futures prices. Estimation of the historical distribution needs to take two practical factors into account. First, one needs to use a time series of data as long as possible to increase precision of estimates. The longer the sample period is, the more efficient estimator we can get. Second, the estimation methodology needs to account for the potential time-varying nature of physical density, especially to allow for the presence of stochastic volatility in the crude oil market as documented by Trolle and Schwartz (2009). Bansal, Kiku, Shaliastovich, and Yaron (2014) further highlight the importance of time-varying volatility when extracting SPDs from economic data.

We first calculate the time series of daily futures returns $\{R_{t,T}\}_{t=1}^{N}$ from 1990 to 2012 using futures prices. It is equivalent to the continuously compounded returns of holding the futures contract to maturity and realizing returns by closing out the position at the maturity date T. At each time t, we normalize the time series of returns with its sample mean $\overline{R}$ and conditional volatility σ_t, the estimation of which is described as below. This gives a time series of return innovation $\{z_{t,T}\}_{t=1}^{N}$, defined as $(\{R_{t,T}\}_{t=1}^{N} - \overline{R})/\sigma_t$. Then, similar to Jackwerth (2000), we estimate the density with a kernel function using the return innovation at t. The physical density of returns is then obtained by $\widehat{P}(R_{t,T}|\mathcal{F}_t) = \widehat{P}(\overline{R} + \sigma_t \cdot z_{t,T})$.

We utilize high-frequency intraday oil futures prices to estimate volatility. Andersen and Bollerslev (1998) and others have shown the superior property of volatility estimated from intraday high frequency data compared to daily data. We calculate daily volatility using the two-scale estimation approach, which Andersen, Bollerslev and Meddahi (2011) have shown to be robust to the impact of microstructure noise in high-frequency data.

In order to match the forward looking horizon of option-implied risk neutral density, we compute expected 21 business-day volatility using the heterogeneous autoregressive (HAR) model proposed by Corsi (2009). We first estimate the regressions of

$$Vol_t = a + b_d Vol_{t-1:t} + b_w Vol_{t-5:t} + b_m Vol_{t-21:t} + e_t,$$

where $Vol_{t-1:t}$, $Vol_{t-5:t}$, and $Vol_{t-21:t}$ denote the most recent daily, weekly, and monthly volatility, respectively. Then we use the HAR regression to predict realized volatility for the next 21 days $\sigma_t = E_t\left[Vol_{t+1:t+h}\right]$, with h equal to 21. The HAR regressions have been used in many studies including Busch, Christensen, and Nielsen (2011) to forecast volatilities in various financial markets. We estimate the HAR regression coefficients on rolling samples of 250 days. This estimation of volatility is free from the look-ahead bias, since at any time t we only use realized historical returns.

Finally, we interpolate the physical density of returns onto the same spacing as the risk-neutral density so that the SPD in the crude oil market is obtained by

$$\widehat{\xi}(R_{t,T}|\mathcal{F}_t) = e^{-r(T-t)} \frac{\widehat{P}^Q(R_{t,T}|\mathcal{F}_t)}{\widehat{P}(R_{t,T}|\mathcal{F}_t)}. \tag{8}$$

Our estimated SPD in the crude oil market is the discounted ratio of option-implied density divided by physical density estimated from futures prices.

2.4 Futures and Option Data

Our crude oil futures and option data are from two sources. We obtain daily crude oil futures and option data from January 2, 1990 to December 31, 2012 from the Chicago Mercantile

Exchange (CME Group, formerly NYMEX); and we get the high-frequency intraday oil futures prices from TickData. Crude oil traded on the CME is the world's largest and most liquid commodity. The range of maturities of futures, the range of maturities and strike prices of options are also larger than other commodity derivatives. An advantage of the data is that crude oil futures and options have been traded on this exchange for more than 20 years, which allows us to study a long time series spanning recessions and many geopolitical events such as the gulf wars, the 9/11 terrorist attacks, the recent financial crisis, and especially the recent boom and bust of the commodity markets. This dataset also provides open interest and trading volume of both options and the underlying futures contracts.[6] Relative demand and trading volume in futures and options reveal investor beliefs and expectations (e.g., Buraschi and Jiltsov, 2006), and therefore are informative when studying the SPDs.

The calculation of the risk-neutral density in (7) is based on European options, but the crude oil option data are American type. We convert American option prices into European option prices following Trolle and Schwartz (2009) who use the methodology of Barone-Adesi and Whaley (1987). After obtaining European option prices, we exclude those observations with Black (1976) implied volatility less than 1% or greater than 200%; we exclude those options with prices less than $0.05 and contracts violating standard no-arbitrage constraints. The empirical analysis is at the weekly frequency and uses OTM calls and puts. Using OTM options is due to two motivations: First it minimizes the effect of possible approximation errors in the early exercise premium; and second, OTM options are usually more liquid than in-the-money (ITM) options. Each week, we use Wednesday data since it is the day least likely to be a holiday during a week. In addition, it is also less likely to be affected by day-of-the-week effects. This selection of data has been widely employed in the literature (e.g., Bates, 1996, 2000; Heston and Nandi, 2000). Since the calculation of risk-neutral density is based on call prices, we utilize OTM puts and transform them into ITM calls. Together with observed OTM calls, call option data effectively span the entire moneyness to apply formula (4).

Panel A of Table 1 provides descriptive statistics for the futures data by maturity. Although the number of contracts and average prices are relatively constant across maturities, average open interest and trading volume decrease sharply beyond the two-month maturity. Open interest of six-month futures contracts is around 20% of one-month contracts, while the trading volume of six-month contracts is only about 4% of one-month contracts. This shows that long maturity futures often lack liquidity, which is also true for options as re-

[6] Crude oil futures (option) trading volume data are missing from December 15, 2006 (December 1, 2006) to May 21, 2007 due to technical reasons when the CME group converted data from the NYMEX database. But option price and open interest data are available throughout the period.

ported in Panel B. Trading volume of all option contracts beyond six-month (or 180 calendar days) is only about 5% of one-month contracts. To ensure our option data have adequate liquidity, we use futures data with maturity of one months and options written on these futures contracts. Panel C of Table 1 reports option data across moneyness. We observe that although deep OTM (ITM) options have large amount of contracts, at-the-money options are most heavily traded. Across moneyness, the average Black (1976) implied volatility displays a smile pattern with deep OTM (ITM) options having higher implied volatility than ATM options. Across maturities, short maturity options on average have a higher implied volatility than long maturity options as shown in Panel B.

2.5 Empirical SPDs in the Crude Oil Market

2.5.1 The U-shaped SPDs

We compute conditional risk-neutral densities and physical densities on each Wednesday using the approach we discussed in section (2.3). While the risk-neutral density can be either negatively or positively skewed, the physical density is relatively more symmetric. The risk-neutral density has fatter tails than the physical density, especially on the negative return side. When we compute the time series of skewness and excess kurtosis, risk-neutral skewness (or kurtosis) dominates physical skewness (or kurtosis); the magnitudes of physical skewness and excess kurtosis are only a small fraction of the ones in risk-neutral distributions. Therefore, the shape of the SPD, which we calculate as the log ratio of the risk-neutral and physical densities, is mainly driven by the risk-neutral density. Figure 1 shows the estimated SPDs as a function of futures returns in the crude oil market on each Wednesday for one-month maturity for the years 1998 to 2012. The horizontal axis denotes returns, and the sample year is indicated in the title of each graph.

The SPDs are nonlinear and in general display an asymmetric U-shape as a function of returns. At the aggregate level, investors in the crude oil market regard the states with extremely low returns or extremely high returns as bad states and assign a high value for payoffs in those states. This might be due to the heterogeneous nature of investors in the crude oil derivative market: Investors (such as long speculators) who have net long futures positions, will bear losses in the case of futures price decreasing if their positions are not protected. They regard negative returns as bad states and highly value an one-dollar payoff in these states. Investors (such as short speculators) who hold net short futures positions, will suffer from increasing futures price and consider those states with positive returns as bad states. They assign a high value for an one-dollar payoff in states with large positive returns.

The U-shaped SPDs display dramatic variations across time. First, we observe that across time investors assign different values of state price to the same level of returns. For a given level of returns, the state price density is higher in 2000 than in 2005, which means investors could have priced a higher payoff for exposure of the same negative or positive returns in 2000 than in 2005. The asymmetric U-shape is wider in the year of 2008-2009 than in other years. Second, the slope of the U-shape has different level of dispersion in different years. For example, consider the year of 2000 and 2007: while investors assign similar state price to certain returns in 2000, their value for the same level of returns varies much more in 2007. Third, the dispersion level of the slopes in both the decreasing and the increasing regions starts to increase around the year of 2004 or 2005. The slopes in both the decreasing and increasing regions can become steeper or flatter depending on the state of nature. In other words, investors' marginal value of one-dollar payoff in difference states, when returns are negative or positive, exhibits significant changes across time. We next analyze returns of OTM puts and OTM calls, i.e. those assets pay off when futures have large negative and positive returns, and we provide further evidence supporting the SPDs with both decreasing and increasing regions.

2.5.2 Option Returns Support the U-shaped SPDs

Net long investors in the futures market face the possibility of losses when futures prices decrease. They may pay premia and buy OTM puts to hedge. Similarly, net short investors in the futures market can seek protection by buying OTM calls. Comparing with producers and inventory operators who have physical assets as the natural hedge of the futures market, speculators have more risk exposure to the futures market and have stronger motivation to hedge with options. By checking the U.S. Commodity Futures Trading Commission (CFTC) futures-only and futures-and-options Commitments of Traders (COT) reports, we find that option positions taken by large traders support this argument. It shows that option positions held by speculators (or non-commercial investors who are not directly engaged in crude oil business activities and use derivatives markets for financial profit purposes) increase faster than those held by hedgers (commercials investors who have direct exposure to the underlying crude oil commodity). Especially after the financialization of commodity markets, the proportion of option positions from speculators grows much faster than the proportion from hedgers.

The premia paid by investors for hedging purposes will increase OTM option prices and cause negative expected returns. In particular, returns of OTM puts will be negative and will increase with strike prices since deep OTM puts (with low strike prices) provide protection for extremely bad states for long investors; returns of OTM calls are negative and will decrease with strike prices since deep OTM calls (with high strike price) provide protection for extremely bad states for short position investors.

We compute option returns as follows. On the third Wednesday of every month, we take a long position in available calls and puts with maturity as close to 30 calendar days as possible. We compute the hold-to-maturity returns of calls and puts as

$$\begin{aligned} r_{t,T}^{call} &= max(F_{t,T}e^{R_{t,T}} - K, 0)/C_{t,T} - 1, \\ r_{t,T}^{put} &= max(K - F_{t,T}e^{R_{t,T}}, 0)/P_{t,T} - 1, \end{aligned} \tag{9}$$

where $F_{t,T}$ is the price of underlying futures, $R_{t,T} = \log(F_T/F_{t,T})$ is the return of underlying futures contract over the period $T - t$, K is the strike price, $C_{t,T}$ and $P_{t,T}$ are prices of European style call and put options. We use observed option prices and we do not create artificial prices by interpolation or extrapolation. For each Wednesday, we assign available options to various bins according to their moneyness defined as $X = K/F_{t,T}$, and returns are averaged within each bin. This provides a non-overlapping return time series with various moneyness. For each moneyness, we show in Table 2 the average return, its standard t-statistics, the 10^{th}, 50^{th} and 90^{th} percentile from January 1990 to December 2012, as well as average returns for two sub-sample periods: pre-financialization January 1990 to December 2004, and post-financialization January 2005 to December 2012.[7] We also investigate oil option returns before and after the bust of oil futures market ($2005-2008$ vs. $2009-2012$) as a dichotomy analysis. We consider five bins of moneyness: $(0.5, 0.85]$, $(0.85, 0.95]$, $(0.95, 1.05]$, $(1.05, 1.15]$ and $(1.15, 1.5)$. OTM puts have moneyness of $(0.5, 0.85]$ and $(0.85, 0.95]$; OTM calls have moneyness of $(1.05, 1.15]$ and $(1.15, 1.5)$. We are mostly interested in returns of deep OTM puts ($X \in (0.5, 0.85]$) and deep OTM calls ($X \in (1.15, 1.5)$).

Panel A of Table 2 reports put option returns. For the period of 1990 to 2012, put options mostly have negative returns and returns increase across moneyness. Negative returns of OTM puts have higher magnitude than ITM puts and are statistically significant. When we compare the difference between OTM puts and ITM puts, returns from OTM puts are statistically lower and more negative. Call option returns are somewhat different from put returns as shown in Panel B of Table 2. There is no clear monotonic pattern of returns across moneyness. Although ITM and ATM call options can have positive returns, returns of deep OTM calls are always negative and statistically significant. In addition, returns from OTM calls are always lower than returns from ITM calls. It means that both deep OTM puts and

[7] Our results remain if we directly work on American option prices or if we date the financialization back to 2004.

deep OTM calls are more expensive relative to other options.

When we compare the pre- and post-2005 periods, we find that the negative return pattern of put options becomes weaker in the post-2005 period than in the pre-2005 period, since only deep OTM puts have significantly negative returns after 2005. On the other hand, OTM call options become more expensive after 2005 as returns of OTM calls are more negative. This could be because short position speculators (who do not have natural hedge and use options to hedge their futures positions) have exposure to the futures price risk, and demand OTM call options to hedge.[8] It further raises OTM call option prices and decreases returns of OTM calls. When we compare the period before and after the bust of oil futures market, we find interestingly OTM put options are less expensive and investors pay less premia to hedge negative returns during $2005-2008$. After 2008, both OTM puts and calls become expensive and have negative returns again.

We further conducted two more analyses. The first analysis is to check whether the expensiveness of OTM calls (and puts) is due to illiquidity premia. We do not find supporting evidence for this. In our sample period, both trading volume and the ratio of trading volume over open interests of OTM calls (puts) are higher than ITM calls (puts), and OTM calls and puts are more liquid than ITM calls and puts. Therefore, OTM calls and puts are actively traded, and the negative returns cannot be imputed to illiquidity. The second analysis is return of the butterfly spread. On the third Wednesday of every month, we take a long position in the butterfly spread constructed from call options with maturity close to 30 calendar days. We buy an ITM call option and an OTM call option with strike prices K_1 and K_3, and we sell two ATM call options with a strike price of K_2 and $K_2 - K_1 = K_3 - K_2$. We find that average return of the butterfly spread is negative when K_3 is above certain threshold. It is consistent with the above argument that OTM call options are expensive relative to other calls, which supports the increasing SPDs in the region of large positive returns.

In summary, we find that put option returns are negative and increase across the strike price. Call option returns are negative and decrease across strike prices when strike prices are high enough, especially after 2005. Negative returns of OTM puts and calls lend support to the U-shaped SPDs; and the different pattern of put returns and call returns supports the asymmetry of the U-shaped SPDs. In addition, option returns are affected by speculations in the oil futures market. Our evidence of negative returns on OTM call options is consistent with the model of Bakshi, Madan and Panayotov (2010) where investors have heterogeneous

[8]In general speculators have net long positions in aggregate to provide liquidity for commodity producers (hedgers). However, there are certain number of short position speculators, such as the Morningstar Short/Flat Commodity Index and the Morningstar Short-Only Commodity Index. In addition, some excessive long positions of financial traders need short positions from other traders, who may hedge themselves.

beliefs and can take short positions.

2.6 Option Implied Risk-Neutral Moments

We are not only interested in the empirical shape of SPD in the crude oil market, but also how SPDs are affected by investors' heterogeneous beliefs. This is related to the economic question of how much more investors in the aggregate are willing to pay for securities in one state over another. Since the historical distribution of futures returns is approximately symmetric, the shape of the SPDs is mainly driven by the properties of the risk-neutral distribution. We compute risk-neutral moments and link their time variations to investor beliefs and trading activities.[9] Risk-neutral moments are calculated from the estimated risk-neutral densities:

$$Var_{t,T} = E_t^Q\left[\left(R_{t,T} - E_t^Q[R_{t,T}]\right)^2\right] \tag{10}$$

$$Skew_{t,T} = \frac{E_t^Q\left[\left(R_{t,T} - E_t^Q[R_{t,T}]\right)^3\right]}{\left\{E_t^Q\left[\left(R_{t,T} - E_t^Q[R_{t,T}]\right)^2\right]\right\}^{3/2}} \tag{11}$$

where $E_t^Q[x] = \int_{-\infty}^{\infty} xP^Q(x)dx$ is the expected value under the risk-neutral measure and $P^Q(x)$ is the density estimated from option prices. Since risk-neutral moments are calculated using the option implied densities and available returns on each Wednesday, they are model free and conditional.

Figure 2 displays the time series of weekly option-implied variance (the upper panel) and skewness (the lower panel) from 1990 to 2012. From left to right, the first two vertical dotted lines denote two significant days of the first Gulf War: the Iraq invasion of Kuwait on August 2, 1990, and the liberation of Kuwait on January 17, 1991. Other vertical dotted lines denote the September 11, 2001 terrorist attacks; the second Gulf War on March 20, 2003; the week when the WTI crude oil spot price reached the highest level in history (July 23, 2008 [$145.31]); the week when the Lehman Brothers filed for Chapter 11 bankruptcy protection (September 15, 2008); the week when the crude oil spot price reached its lowest level during the recent crisis period (December 23, 2008 [$30.28]); the week when the Libyan Civil War began and oil and gas production in Libya fell by $60\% - 90\%$; and the week when Standard & Poor's downgraded the U.S. long-term sovereign credit rating from AAA to AA+,

[9] Datta, Londono and Ross (2014) investigate how option-implied moments change around important events in the crude oil market.

respectively. We note that variance rises sharply on those days. Consistent with Robe and Wallen (2014), these sharp hikes in the upper panel of Figure 2 indicate that oil variance is affected by not only the oil market fundamentals, but also macroeconomic conditions. However, when we compare the upper and lower panels of Figure 2, it appears that skewness reacts more to oil market-specific information. For example skewness significantly increases during the week of the second Gulf War in March 2003 and the week of Libyan War in February 2011, but it does not change much when the U.S. sovereign credit is downgraded.

In summary, the empirical results show that the SPDs in the crude oil market display an asymmetric U-shape as a function of returns, and exhibit remarkable variations across time. Returns of OTM options are consistent with this nonlinear pattern of the SPDs. We also find some preliminary evidence that risk-neural skewness of oil futures returns is more likely to be affected by oil market-specific information.

3 Investor Beliefs and the SPDs

The literature has shown that the SPDs depend on investor disagreements. Heterogeneity is represented not only in asset prices and returns, but also in the relative positions and trading volumes in equilibrium. Dependence of the SPDs on heterogeneous beliefs is present in the equity market (e.g., Anderson, Ghysels and Juergens, 2005; Buraschi and Jiltsov, 2006), as well as in the foreign currency market (Beber, Buraschi and Breedon, 2010). When agents have different beliefs about future price movements, they engage in trading either for speculation or hedging.

How do investor beliefs affect the SPDs in the crude oil market? In this section, we investigate how investor beliefs embedded in crude oil derivatives trading affect the slope of the SPDs, because it is one of the fundamental characteristics of the SPDs and can be directly related to investors' risk aversion. The steeper the slope of SPDs is, the higher state prices investors assign at certain economic state. We first describe how investors' trading positions have evolved over the past twenty years in crude oil futures and options. We then discuss the measures of investor beliefs in both the crude oil futures and option markets. Subsequently, we document that the slopes of the SPDs in both decreasing and increasing regions are affected by investor beliefs. We further substitute the slope with risk-neutral skewness and investigate the impact of heterogeneous beliefs.

3.1 Market Participation and Measure of Beliefs

The CFTC classifies large traders in the crude oil derivatives market into commercial traders or hedgers and non-commercial traders or speculators.[10] Figure 3 shows long and short positions taken by hedgers and speculators in the futures market as well as their ratios, which we obtain from the CFTC futures-only Commitments of Traders (COT) report. Although participation in the futures market by hedgers and speculators has experienced steady growth from 1990 onwards, the increases in positions have been faster since 2004-2005. When we look at the ratio of long positions taken by speculators over hedgers, we see the ratio typically fluctuates around 0.2 before 2004-2005, but it has consistent growth afterwards. The ratio of short positions also has steady growth, although not as substantial as the ratio of long positions. The fast growth of positions from speculators and the historical rise of futures price in 2008 and the followed sudden drop (as shown in the bottom panel of Figure 3) has drawn significant attention from the academicians, practitioners and policy makers. Some literature has attributed position growth of speculators to commodity linked investments.

The crude oil option market has experienced even more dramatic growth. Figure 4 shows the number of OTM options and their open interests and trading volumes aggregated within each week. OTM calls are defined as call options with moneyness (K/F)>1.05, and OTM puts are put options with moneyness (K/F)<0.95. We only consider options with maturity less than 180 calendar days. We also report the weekly average annualized 30 calendar-day fixed maturity oil option-implied volatility which is constructed using the cross-section of OTM option prices and the methodology of CBOE's VIX. While there are an increasing number of OTM options traded in the market, the growth of the number of options after $2004-2005$, especially the first half of 2008 is indeed dramatic. In particular, the number of OTM call option contracts jumps in 2008. When oil prices are most volatile, and investors (especially short investors) demand OTM options (especially OTM calls) to hedge their positions. We observe that open interests of options have certain decrease after 2010. But trading volume of both OTM calls and OTM puts has remained at a high level. Given the active participation in the futures and option markets, we consider measures of investor beliefs in both markets and investigate how they affect the SPDs. We will focus on those variables measuring the level of belief heterogeneity rather than those variables directly reflect either net long or short positions.

First we use the speculation index, which quantifies the level of speculation activity, as a

[10] We acknowledge that this classification of hedgers and speculators is not perfect and does not cover all investors in the crude oil derivatives market. However, Büyükşahin and Robe (2014) confirm that the speculative activities inferred from the public CFTC position data can capture speculative activities inferred from non-public trader-level CFTC data well.

measure of investor beliefs in the crude oil market. Several studies such as Scheinkman and Xiong (2003) and Xiong and Yan (2010) argue that agents with heterogeneous beliefs engage in speculative trading among each other. The speculation index is to gauge intensity of speculation relative to hedging (Working, 1960; Büyükşahin and Harris, 2011). If we denote SS (SL) as the speculative short (long) position, and HS (HL) as the hedging short (long) position, we define

$$\text{Speculation Index}_t = \begin{cases} 1 + \frac{SS_t}{HL_t + HS_t} \text{ if } HS_t \geq HL_t, \\ 1 + \frac{SL_t}{HL_t + HS_t} \text{ if } HS_t < HL_t. \end{cases} \tag{12}$$

Since speculators take positions in crude oil futures by anticipating certain price movements, this speculation index contains information on belief differences among investors.[11] The speculation index has to be interpreted together with hedging activity in the futures market. It measures the extent by which speculative positions exceed the necessary level to offset hedging positions. Panel A of Figure 5 plots the speculation index from 1990 to 2012. We observe that there has been a high level of speculative activities in recent years. Before 2000s, the speculation index is below 1.05 meaning less than 5% of excessive speculation; however, it rises steadily over time to 1.2 in 2010 and drops to 1.1 afterwards. As such, the premia that hedgers pay for insurance against futures price risk are highly affected by the active participation of speculators.

Besides positions taken by speculators, actual trading volume of futures may reflect the degree of heterogeneity and how investors speculate and share risks among each other. Literature (e.g., Kandel and Pearson, 1995) has documented the positive relationship between investor heterogeneous beliefs and volume of trade. Buraschi and Jiltsov (2006) show that, the trading volume of stocks and options is positively correlated with the dispersion in beliefs. Carlin, Longstaff, and Matoba (2014) show that increased disagreement is associated with higher trading volume. Therefore, we use the trading volume of futures as another measure of investor belief dispersion in the underlying futures market. Panel B plots the trading volume of one-month futures contracts, which is relatively stable before 2004-2005 and significantly magnifies afterwards.

Next we discuss two measures of heterogeneous beliefs in the option market used by Han (2008). One is the open interest ratio of OTM puts to calls, which measures the relative demand for insurance against downside risks and reflects the hedging needs of heterogeneous agents; the other is the trading volume of options, which is a proxy for the level of disagreement among investors. Open interests and trading volume of options do not capture

[11] Singleton (2014) discusses how disagreements among investors induce speculative activities, price drift, and high volatility in the crude oil market.

the same information since open interests are the outstanding positions investors take, while trading volume can be due to opening or closing a contract. A high open interest ratio of OTM puts to calls suggests investors are overall more pessimistic, they tend to buy put options either to get protection against future price drops or to pursue potential returns on put positions. Panel C shows that the open interest ratio of OTM puts to calls written on the one-month futures contracts. It shows several spikes during our sample period. Panel D shows that, the option trading volume starts to rise steadily from 2004-2005 and becomes much more volatile afterwards. In the subsequent regression analysis, we adjust the time series of futures trading volume and option trading volume with a deterministic time trend following the literature.

To what extent do investor beliefs of the economy and the equity market affect the SPDs in the crude oil market? We consider two types of proxies of beliefs to address this question: investor beliefs about the equity market and economy, as well as investors' expectations about the market volatility. The first measure is the bull-bear spread based on the survey of Investors Intelligence which has been used by Brow and Cliff (2004, 2005) and Han (2008). Every week, Investors Intelligence sends out 150 surveys to institutional investment advisors and collects their expectations of future market movements as bullish, bearish, or neutral. The bull-bear spread is then calculated as the percentage of bullish investors minus the percentage of bearish investors, and it is often used as a proxy for beliefs of institutional investors. The second type of proxy we use to measure investor beliefs about the stock market is the CBOE's VIX index. The VIX has become a benchmark for measuring investors' expectations of market volatility and investor sentiment as a fear indicator. As plotted in Panels E and F, the VIX is in general negatively correlated with the bull-bear spread. The VIX tends to be higher when more investors are less confident of the market. Investors are likely to trade more when there is higher uncertainty. We also consider the consumer sentiment index from the University of Michigan and the investor sentiment index in Baker and Wurgler (2006) as plotted in Panels G and H. Literature has used the Michigan sentiment index to study how sentiment affects stock prices (e.g, Lemmon and Portniaquina, 2006; Stambaugh, Yu, and Yuan, 2012).

3.2 Other Control Variables

We consider two sets of control variables in the regression analysis. One set includes variables specific to the crude oil market that determine futures returns and therefore may affect the

SPDs. We first follow Hong and Yogo (2012) and calculate the basis as

$$\text{Basis}_{t,T} = (\frac{F_{t,T}}{S_t})^{\frac{1}{T-t}} - 1, \tag{13}$$

where $F_{t,T}$ is the price of futures with maturity of $T - t$ at time t, S_t is the spot price at time t. Basis is the spread between futures prices and spot prices, and it can be interpreted as the implied net convenience yield.

Next we include option implied volatility due to the following considerations. First, Trolle and Schwartz (2009) present strong evidence of stochastic volatility in the crude oil market; and stochastic volatility has to be incorporated into the model to price commodity derivatives. Therefore, volatility could affect futures prices and returns. Second, contrary to the traditional theory of storage, which claims volatility tends to be high when the futures price is in backwardation, Carlson, Khokher, and Titman (2007) and Kogan, Livdan, and Yaron (2008) document that the relationship between the futures term structure slope and volatility is non-monotonic. While the sign of the slope of the futures term structure implies investors' demand for the convenience yields, volatility captures complementary information, indicating investors' exposure to risk.

We also consider the storage level of crude oil and historical returns of futures, both of which affect futures returns as documented in the literature (e.g., Bessembinder, 1992; Bessembinder and Chan, 1992; de Roon, Nijman, and Veld, 2000). We use the storage level defined as U.S. total stocks of crude oil, excluding strategic petroleum reserves, based on the report from Energy Information Administration (EIA). Following Acharya, Lochstoer, and Ramadorai (2013), a Hodrick-Prescott filter is applied to remove the trend, where the smoothing parameter is set to a number appropriate for weekly data. The historical return is the moving average of daily return of holding a long position of the futures contract over one week, since our regression is weekly.

The other set of control variables capture the macroeconomic condition and investment opportunities. They include 1-month treasury-bill rate, the spread between yields on the 10-year treasury bond and the 3-month treasury bill, the difference in yields on Baa and Aaa corporate bonds, the Chicago Fed National Activity Index and the log growth in aggregate industrial production over the last 12 months. These variables are used by Han (2008) and others. We obtain the one-month treasury bill rate from the online data library of Kenneth French, and other macro variables are from the Federal Reserve Bank of Chicago and St. Louis.

Table 3 reports the summary statistics of risk-neutral skewness computed from (11), main regressors and control variables in the crude oil market. We apply necessary filters to some regressors to remove time trends. The last column of Table 3 show that the unit-root tests are rejected for all time series we use in regressions at the 5% significant level. Note that overall skewness is negative; investors on average demand more OTM puts than OTM calls; and historical annualized return of holding the one-month futures contract is 7.56% with a sharpe ratio of 0.47. Table 4 reports the correlation matrix of regressors, as well as control variables in the crude oil market. We observe that all correlations are rather low, except for volatility having a correlation of 0.65 with VIX, Bull-bear spread having a correlation of -0.28 with VIX, and OTM put/call ratio having a correlation of -0.36 with basis.

3.3 Regression Results

We consider three levels of measures to examine how investor beliefs affect the slope of the SPD (which characterizes marginal utility of investors) in the crude oil market: 1) belief dispersion in the crude oil futures market; 2) belief dispersion in the crude oil option market; and 3) investor beliefs of the equity market and the general economy. In our regression, we add control variables discussed above which may affect crude oil futures returns and reflect macroeconomic conditions. The full regression model is given by

$$\begin{aligned} \text{Slope of the SPD}_t \;=\;& \alpha_t + \beta_t \cdot \text{Beliefs in the crude oil futures market}_t \\ &+ \gamma_t \cdot \text{Beliefs in the crude oil option market}_t \\ &+ \delta_t \cdot \text{Beliefs in the equity market}_t \\ &+ \eta_t \cdot \text{Slope of the SPD}_{t-1} \\ &+ \text{Control variables}_t. \end{aligned} \tag{14}$$

To assess the dependence of the slope of the SPD on investor beliefs, we run the regression (14) using different dependent variables. We first directly measure the slope of the SPD curve based on different return intervals. Second, we substitute risk-neutral skewness for the direct slope measure, since the asymmetric U-shape pattern of the SPDs is mainly determined by the level of risk-neutral skewness. We also compare regression results with risk-neutral moments obtained from another model-free methodology developed by Bakshi, Kapadia, and Madan (2003; BKM hereafter), which uses OTM option prices. We include the lagged dependent variable to control for its positive autocorrelation and we standardize all variables before performing regressions. To minimize the estimation noise due to limited option data, we only include those days with at least two OTM calls and two OTM puts in the regression sample to ensure the estimation of our dependent variable is reliable.

3.3.1 Dependence of the Slope of the SPDs on Investor Beliefs

Since the SPD in the crude oil has a time-varying U-shape, investors assign high state price over large negative and positive returns. The decreasing- (increasing-) region slope of the U-shaped SPDs captures investors' preferences towards different level of negative (positive) returns. The steeper the slope, the higher state price investors assign. On each Wednesday, we calculate the measure of slope as the difference between the SPDs of two return points:

$$\begin{aligned} \text{Decreasing-Region Slope}_t &= \text{SPD}_t(R_2) \text{ - } \text{SPD}_t(R_1), \text{ where } R_2 < R_1 << 0; \\ \text{Increasing-Region Slope}_t &= \text{SPD}_t(R_2) \text{ - } \text{SPD}_t(R_1), \text{ where } R_2 > R_1 >> 0. \end{aligned}$$

This direct measure of slope shares the same idea with Xing, Zhang and Zhao (2010) who quantify the steepness of stock option smirks. We consider two measures of the decreasing-region slope based on return intervals $(-0.20, -0.15)$ and $(-0.18, -0.12)$, and two measures of the increasing-region slope based on $(0.10, 0.16)$ and $(0.13, 0.18)$. In the regression analysis, we include basis, storage level, historical returns, and option implied volatility as control variables and we trim the outliers beyond the 2.5^{th} and 97.5^{th} percentile.

Table 5 presents regression results. It shows that futures trading volume negatively affects both the decreasing- and increasing-region slope, which means state prices for both negative and positive returns are higher when investors in the future market trade less. The open interest ratios of OTM puts over OTM calls (ratios of OTM calls over OTM puts) positively affect the steepness of the decreasing- (increasing-) region slope. In other words, when investors demand more OTM puts (calls), they assign higher state price to negative (positive) returns. Comparing the pre- and the post-financialization periods, speculation has a distinct impact on investors' marginal utility in the two periods. Investors' marginal utility over positive returns significantly decreases given the high level of speculation after $2004-2005$. We also observe that coefficients of investor beliefs of the equity market and the economy are not robust.

Although we present the evidence that the slope of the SPD is affected by investor beliefs in the crude oil market, we have to discuss the decreasing- and increasing-region slope separately. In addition, there are multiple return intervals on each day and different choices may change the magnitude of slope. Therefore, this measure of slope only represents investors' preference over certain returns and can be arbitrary. Next we argue that risk-neutral skewness is a more comprehensive measure of the slope and our later results mainly focus on the risk neutral-skewness.

3.3.2 Option-Implied Skewness and Investor Beliefs

The asymmetric U-shape pattern of the SPDs arises from the differences between risk-neutral skewness and physical skewness. Because the skewness of the risk-neutral density has more pronounced time variation compared with the physical density, the slope of the U-shaped SPDs mostly depends on the skewness of the risk-neutral density. Furthermore, skewness measure utilizes all data available with valid closing quotes. It provides a thorough picture of the shape of the SPDs. Therefore, risk-neutral skewness is a more reliable and comprehensive measure of the slope (or variations in the marginal rate of substitution across states). We next rely on risk-neutral skewness and investigate whether it is affected by investor beliefs. In the following regressions, dependent variables are risk-neutral skewness obtained from (11).

Table 6 contains the basic regression results for three sub-sample periods 1990-2004 (pre-financialization), 2005-2012 (post-financialization) and 2005-2008 (when the oil market is most volatile).[12] The first four columns of each panel show results for 1990-2004. The results suggest that risk-neutral skewness is negatively associated with the OTM puts to calls open interest ratio and positively associated with belief dispersion in the option market measured by options volume. A high open interest ratio of OTM puts to calls suggests investors are overall more concerned about potential large negative returns, and is therefore associated with high state prices of large negative returns. When there is more option trading volume, risk-neutral skewness tends to be higher and investors overall assign higher state price over large positive returns. Comparing the pattern of coefficients for the pre- and post- financialization periods in all three panels, the most striking point is that the speculation index becomes a significant factor for negative risk-neutral skewness after 2005. More speculative trading implies a more negatively skewed distribution, especially for the 2005-2008 period. The effect of speculations also amplifies in this period as the magnitude of coefficients strongly increases. It suggests that the increasing speculative activities after 2004-2005 do have a significant impact on the futures market, lending support to the argument of Singleton (2014): more speculation after 2004-2005 may cause the high crude oil price, meaning a high probability of negative expected returns. Furthermore, skewness becomes less persistent after 2005. One possible explanation is that investors start to trade more often and incorporate new information and change their expectations faster than before. Lastly, regression results show that investor beliefs about the equity market do not have a significant impact on the skewness of crude oil futures returns or on the shape of the

[12] When including the investor sentiment index in Baker and Wurgler (2006) our sample is restricted up to 2010 and all results remain unchanged. Therefore we do not report results with this investor sentiment measure.

SPDs.

Two comparisons with the existing literature can be made here. First, the empirical evidence that the OTM puts to calls ratio affecting the risk-neutral skewness aligns with findings in the stock and index option literature (Dennis and Mayhew, 2002; Han, 2008). Since heterogeneous investors have different demands for OTM options due to their various expectations of market fundamentals, pessimists can share risks with others by buying insurance from optimists. The larger the difference in beliefs, the higher the demand for OTM puts which drives up the prices of options with low strike prices, and the distribution is therefore more negatively skewed. Second, consistent with Buraschi and Jiltsov (2006) we confirm that option trading volume can affect the slope of the SPDs as a proxy for dispersion of beliefs. Option trading volume positively affects risk-neutral skewness, meaning higher level of belief dispersion is related with steeper slope in the increasing region or higher state price of positive returns. One possibility is that short position speculators actively engage in trading both futures and options to share risk or speculate among each other. This is also consistent with the expensiveness of OTM call options we document in Table 2 and net short investors use OTM calls to hedge their potential loss caused by positive returns.

3.3.3 Robustness Check

To further illustrate the dependence of risk-neutral skewness on investor beliefs, we perform two types of robustness exercises. The first exercise is to include control variables in the crude oil market which affect futures returns, as well as macroeconomic variables which represent economic conditions and investment opportunities, in the regression model (14). In the other robustness check we replace the dependent variable by another model-free measure of risk-neutral skewness (BKM, 2003), and we check whether the results obtained from density-based skewness hold.

Panel A of Table 7 shows results when adding control variables in the crude oil market, namely, basis, storage level, historical returns, and option implied volatility. For ease of comparison, we only report coefficients of the same variables as in Table 6, i.e., β_t, γ_t, δ_t and η_t in (14). The coefficients of the additional control variables are omitted for brevity. Comparing with Table 6, adding control variables only marginally increase R^2. The main results still hold for both the significance level and the magnitude of coefficients of interest. Higher demand for OTM puts is associated with more negative skewness. Higher belief dispersion in the option market is associated with more positive skewness. More speculative positions after 2005 induce a higher probability of negative returns and this effect is particularly strong during the period of $2005 - 2008$. It confirms our findings in Table 6 that investors beliefs do affect risk-neutral skewness and state prices of the crude oil market. In addition, belief measures in the equity market do not have significant impacts.

Panel B of Table 7 reports regression results when adding the macro control variables, which are 1-month treasury-bill rate, the spread between yields on the 10-year treasury bond and the 3-month treasury bill, the difference in yields on Baa and Aaa corporate bonds, the Chicago Fed National Activity Index and the log growth in aggregate industrial production over the last 12 months. These variables reflect the general economic condition and may affect the demand and supply of oil futures and therefore their expected returns. Again, we observe that the risk-neutral skewness is more negative (marginal value of a dollar over negative returns are higher) when there are higher OTM put open interest over OTM calls; marginal value of a dollar over positive returns are higher when there are higher level of belief dispersion in the option market; and a higher level of speculative activities after 2005 causes a higher probability of negative expected returns. Belief measures in the equity market have no significant impacts on the SPDs in the crude oil market.

We next use risk-neutral skewness computed from OTM option prices (BKM, 2003). Our numerical implementation of extracting risk-neutral moments follows Duan and Wei (2009). While our calculation of risk-neutral skewness in (2.6) is based on the densities inferred from the cross-section of option prices, risk-neutral skewness of BKM (2003) is calculated from OTM option prices. Since these two methodologies of calculating moments are very different by nature, one cannot expect estimated moments to perfectly accord with each other. For example while the two measures of risk-neutral variance have a correlation of about 95%; the two measures of skewness have a correlation of 70.04% after 2005.

Table 8 reports the regression results using the alternative measure of risk-neutral skewness. All regressions have the same model specification as the regressions in Table 7. The main conclusions still remain valid: Higher OTM put demand implies more negatively skewed risk-neutral densities. A higher level of speculation after $2004-2005$ is associated with more negative skewness. However, we do not find significant relationship between option trading volume and BKM skewness. It is possible that our density-based skewness is calculated from the entire cross-section of option prices and therefore can keenly capture the relationship with trading volume. In unreported results, we run all regression analyses in Tables 6-8 at the monthly frequency. Our main conclusions remain.

In summary, regression results indicate that in aggregate investors assign higher state prices in the crude oil market (i.e., the decreasing-region slope of the SPDs is steeper) when there are higher OTM put demand. More speculative positions after $2004-2005$ induce more negative risk-neutral skewness and reinforce this effect. Higher level of belief dispersion in the option market is associated with higher state prices in the increasing region. This implies that the relative state price change with respect to positive returns becomes higher when investors overall trade more options. Investors' sentiment of the equity market such as bull-bear spread does not have a strong impact on the SPDs in the crude oil market. The findings are consistent and robust, regardless of whether we consider more control variables or use another measure of risk-neutral skewness.

4 Conclusion

Using more than twenty years of futures and option data, we back out investor preferences in the crude oil market. We first estimate return distributions under the physical and risk-neutral measures, and we characterize the time variation in the SPDs. Moreover, we investigate the dependence of the SPDs on investors' heterogeneous beliefs by utilizing the informative features of the crude oil derivatives market. Comparing with other financial markets, speculation data is available in this market and investors can take long or short positions based on their beliefs without being subject to regulations. We investigate how our measures of investor beliefs affect the slope of the SPDs, which can be directly related to investors' preference over large negative or positive returns. We analyze the pre-financialization period from 1990 to 2004, as well as the post-financialization period from 2005 to 2012, when the level of speculative activities in the crude oil market is believed to be high.

We obtain three main results. First, we find that the SPDs in the crude oil market display a time varying U-shape pattern. This implies that investors with various hedging or speculative purposes regard large negative or positive returns as bad states. We show that returns on OTM calls and puts are negative and are consistent with the U-shaped SPDs. Second, we document the dependence of the SPDs on investor beliefs. When there are higher demands for OTM puts, the decreasing-region slope of the SPDs is steeper, and investors assign higher state prices to negative returns. The slope of the SPDs is also affected by belief dispersions measured by options trading volume. These findings are robust under different model specifications. Third, the increase in speculation after 2005 is associated with more negative risk-neutral skewness and higher marginal value of one dollar paid in negative return states.

The empirical findings in this paper suggest several extensions in the broader context of commodity markets. First, to price commodity derivatives, a model needs to have the SPDs that reconcile the deviation between physical and risk-neutral densities. Second, it would be interesting to develop theoretical commodity pricing models with heterogeneous beliefs. Prices are determined in an equilibrium where investors share risks and speculate based on their beliefs about the market. Sockin and Xiong (2014) spearhead this literature by analyzing how informational frictions affect commodity prices. Finally, our results highlight the importance of exploring the implications of heterogeneous beliefs for the relative demand and trading volume of commodity futures and options. When agents have different beliefs of commodity markets, they engage in trading derivatives either for hedging or speculation.

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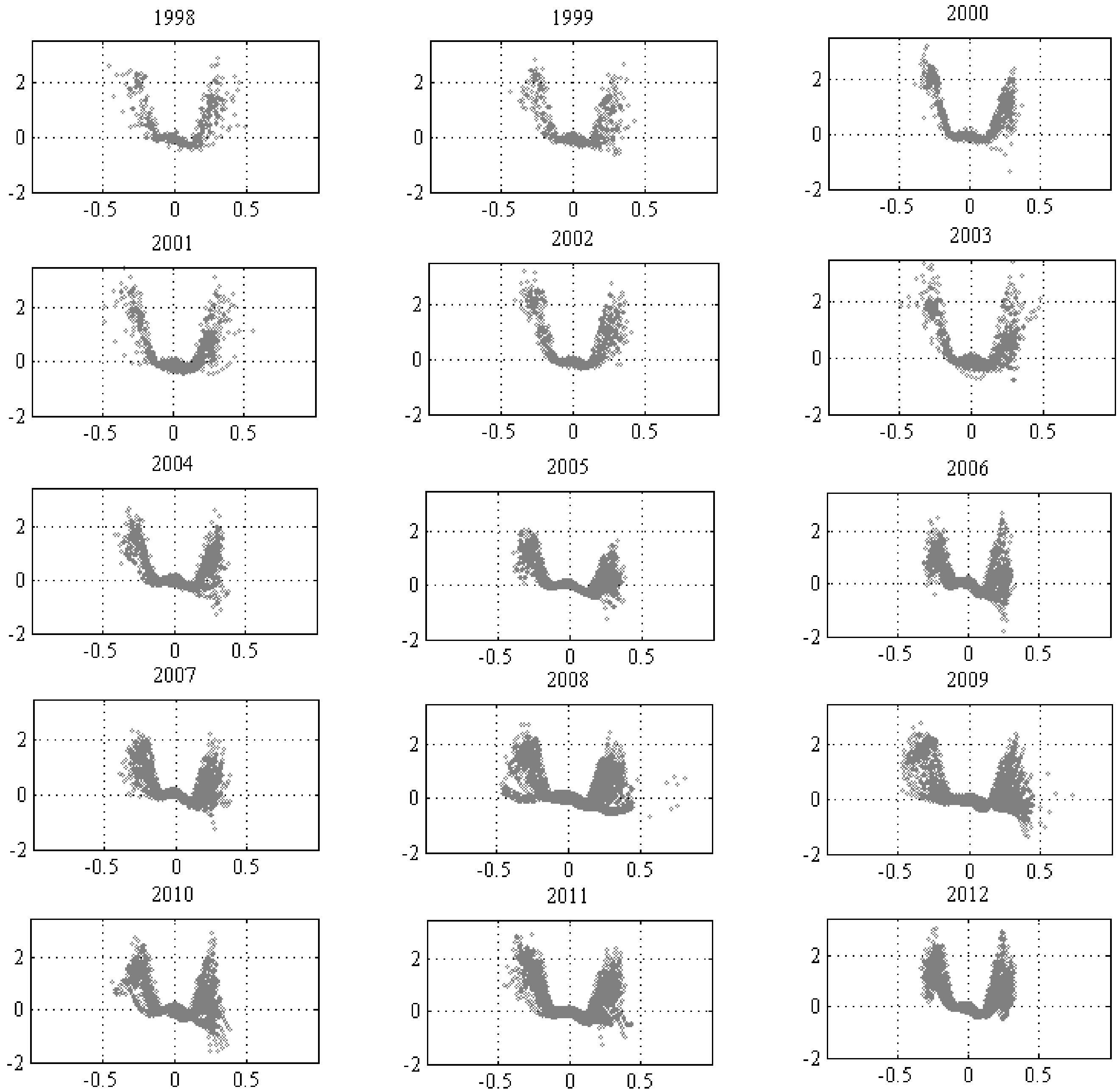
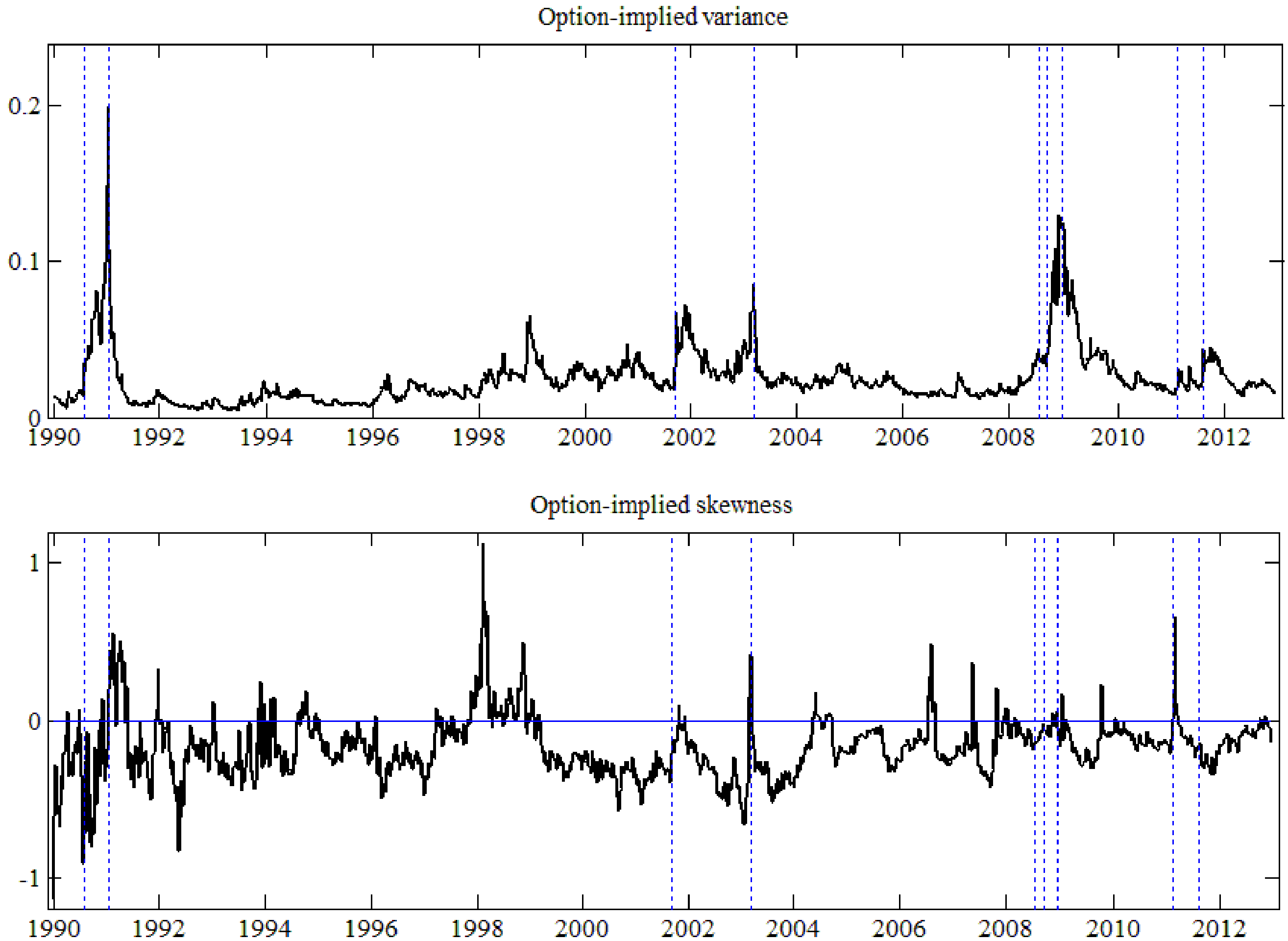
Figure 1: State Price Densities in the Crude Oil Market



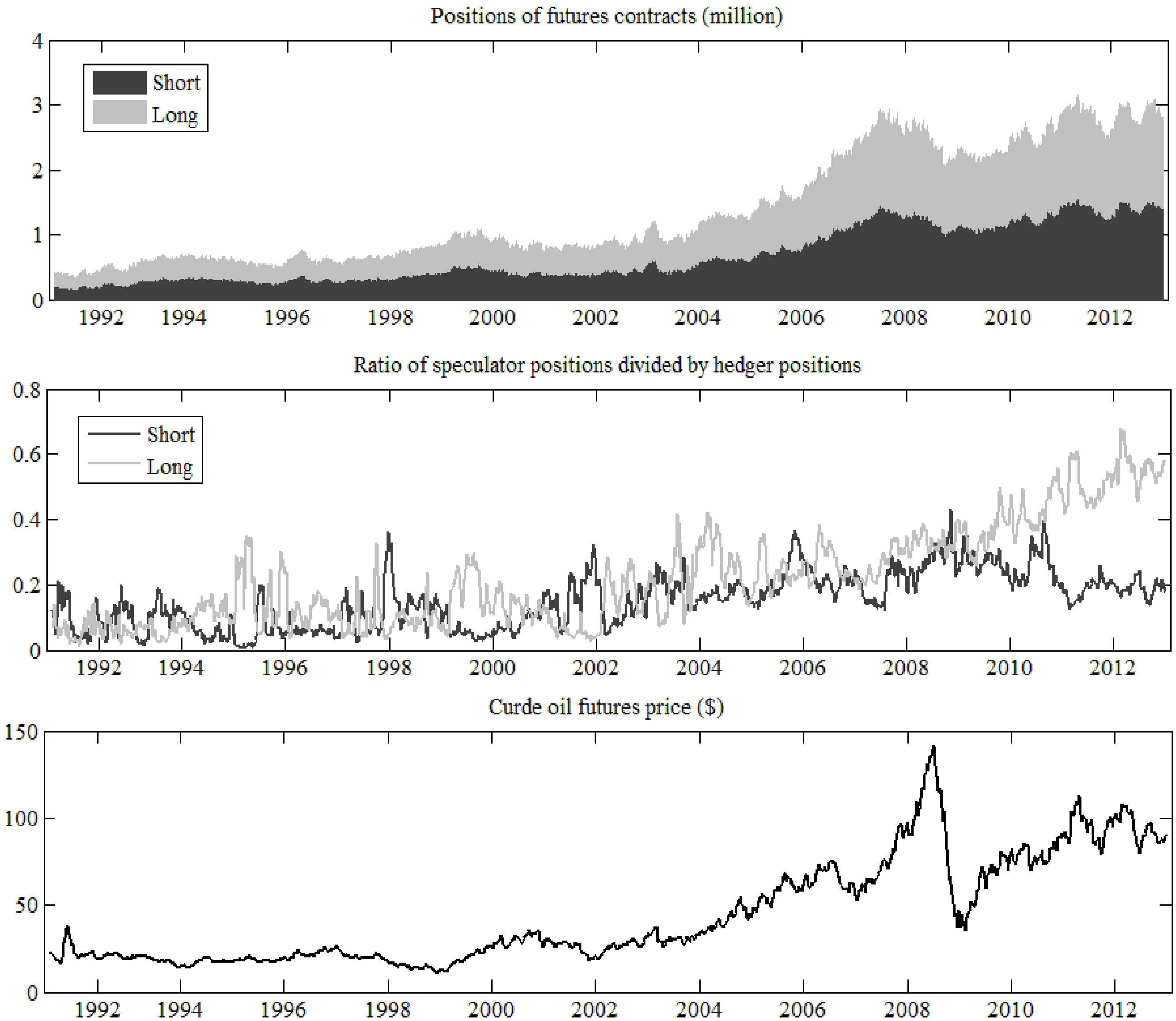
Notes: This figure shows the estimates of the state price densities (SPDs) in the crude oil market in one-month horizon. For each year from 1998 to 2012 we plot the log ratio of the risk-neutral densities and physical densities of futures returns on each Wednesday. The horizontal axis is futures returns defined as $\log(F_T/F_{t,T})$.

Figure 2: Option-Implied Variance and Skewness



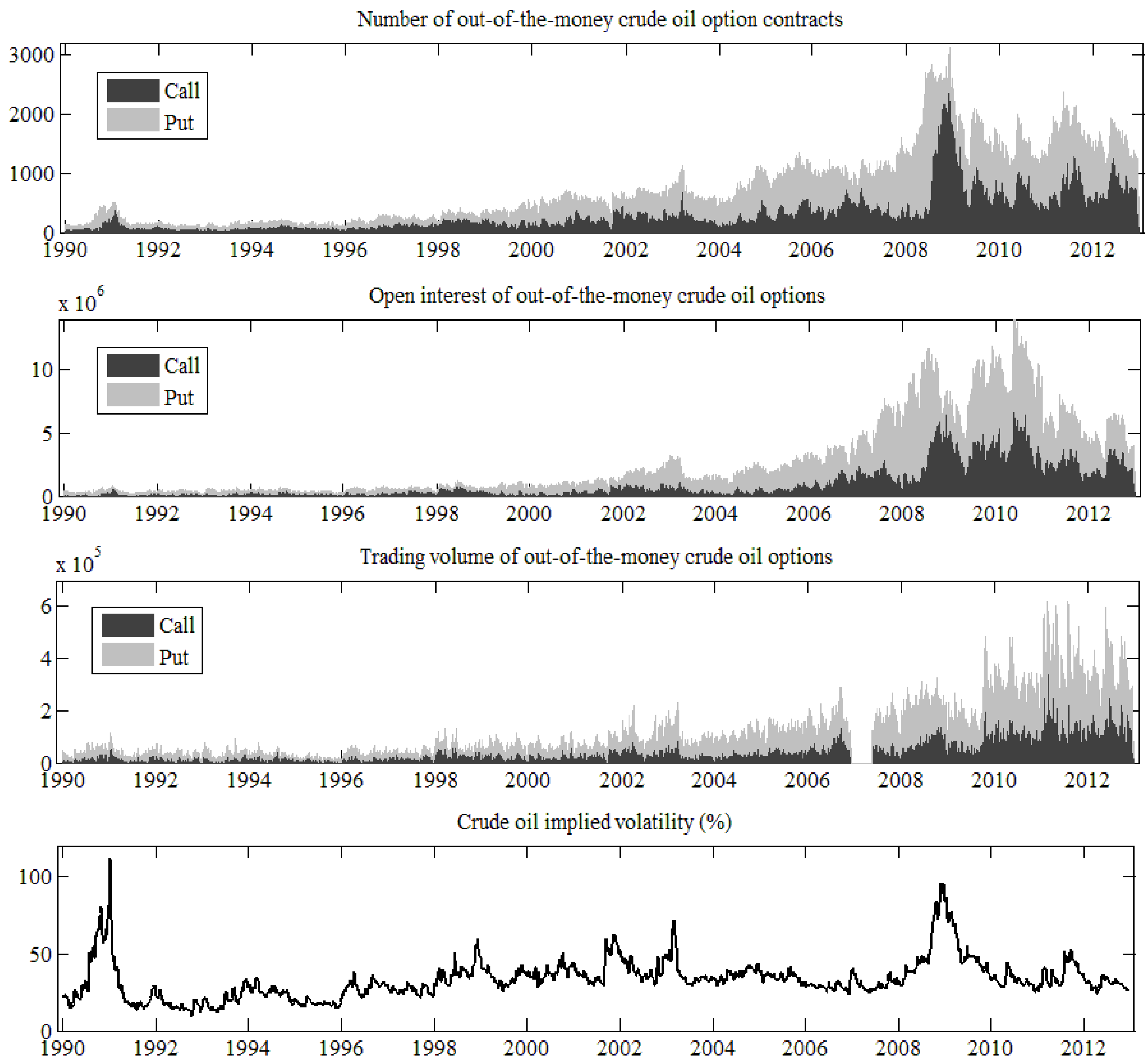
Notes: This figure reports option-implied variance and skewness of crude oil futures returns in one-month horizon on each Wednesday from January 1990 to December 2012. We first obtain risk-neutral densities on each Wednesday using the method of Ait-Sahalia and Lo (1998, 2000). Variance and skewness are then calculated based on estimated densities. From left to right, the first two vertical dotted lines denote two significant weeks of the first Gulf War: the Iraq invasion of Kuwait on August 2, 1990, and the liberation of Kuwait on January 17, 1991. Other vertical dotted lines denote the September 11, 2001 terrorist attacks; the second Gulf War on March 20, 2003; the week when the WTI crude oil spot price reached the highest level in history (July 23, 2008 [$145.31]); the week when the Lehman Brothers filed for Chapter 11 bankruptcy protection (September 15, 2008); the week when the WTI crude oil spot price dropped to its lowest level during the recent crisis periods (December 23, 2008 [$30.28]); February 15, 2011, when the Libyan Civil War began and oil and gas production in Libya fell by 60%-90%; Standard & Poor's downgraded the U.S. long-term sovereign credit rating from AAA to AA+, respectively.

Figure 3: Crude Oil Futures Positions Taken by Hedgers and Speculators



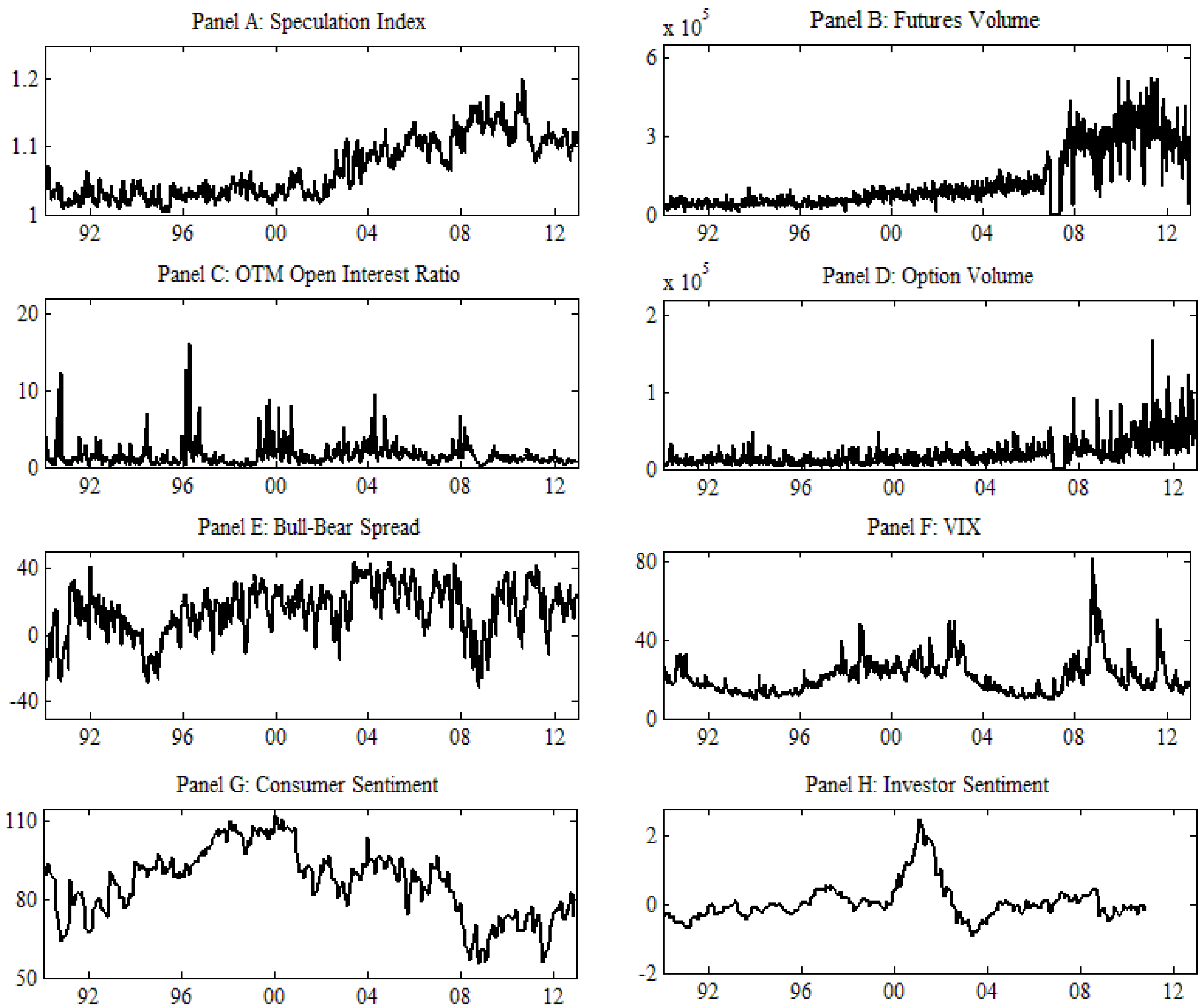
Notes: This figure shows long and short futures positions taken by hedgers and speculators from 1990 to 2012. Positions of hedgers and speculators are from the U.S. Commodity Futures Trading Commission (CFTC) futures-only Commitments of Traders (COT) report. Before September 30, 1992, the CFTC publishes reports twice every month; thereafter, reports are available every week. The top panel displays the aggregate long (short) futures positions taken by large traders; the mid panel shows the ratio of speculator positions over hedger positions; the bottom panel reports the closest-to-maturity futures price averaged within the CFTC reporting week. The full sample covers January 2, 1990 to December 31, 2012.

Figure 4: Number of Contracts, Open Interests, Trading Volumes of Crude Oil Futures Option and Implied Volatility



Notes: This figure reports the number of out-of-the-money (OTM) crude oil option contracts, their open interests and trading volumes aggregated within each week from 1990 to 2012. While OTM calls are defined as call options with moneyness (K/F) >1.05, OTM puts are put options with moneyness (K/F)<0.95. We report all options with maturity less than 180 calendar days. Options trading volume data are not available from December 1, 2006 to May 21, 2007 when the CME group converted data from the NYMEX database. But option price and open interest data are available throughout the period. We also report annualized oil option implied volatility averaged within each week in the bottom panel. The full sample covers January 2, 1990 to December 31, 2012.

Figure 5: Measures of Investor Beliefs



Notes: This figure shows measures of investor beliefs of the crude oil market, the equity market and the economy. Speculation index measures the extent by which speculative positions exceed the necessary level to offset hedging positions. Data is collected from CFTC's trader report. Futures volume is based on the one-month futures contracts. OTM open interest ratio is open interest of OTM puts divided by open interest of OTM calls. OTM open interest ratio and option volume are based on options written on the one-month futures contracts. Bull-Bear spread is defined as the percentage of bullish investors minus the percentage of bearish investors based on the survey of Investors Intelligence. VIX is the CBOE's Volatility Index. Consumer sentiment index is from Surveys of Consumers from the University of Michigan. Investor sentiment is the sentiment index in Baker and Wurgler (2006).

Table 1: Crude Oil Futures and Options Data

Panel A: Futures Data by Maturity (months)

	1	2	3	6	12	24	All
Num. of Contracts	5759	5778	5780	5781	5774	4414	33286
Average Price	42.673	42.773	42.833	42.827	42.527	44.669	42.984
Average OI	141954	120005	62132	28914	15209	5701	64596
Average Volume	113774	62385	22319	4790	1316	317	35492
Average Volume/OI Ratio	0.80	0.52	0.36	0.17	0.09	0.06	0.55

Panel B: Option Data by Maturity (calendar days)

	(0,30]	(30,60]	(60,90]	(90,120]	(120,180]	(180,∞)	All
Num. of Contracts	57032	69285	70475	65641	100265	398810	761508
Average IV	0.459	0.415	0.395	0.380	0.360	0.303	0.348
Average Price	5.799	5.532	5.690	6.127	6.334	10.768	8.466
Average OI	4207	3579	2919	2585	2197	1672	2299
Average Volume	390	207	105	70	48	20	81
Average Volume/OI Ratio	0.09	0.06	0.04	0.03	0.02	0.01	0.04

Panel C: Option Data by Moneyness (K/F)

	(0,0.85]	(0.85,0.90]	(0.90,0.95]	(0.95,1.05]	(1.05,1.10]	(1.10,1.15]	(1.15,∞)	All
Num. of Contracts	151885	62582	76275	165759	56497	43542	204968	761508
Average IV	0.369	0.329	0.316	0.307	0.317	0.327	0.395	0.348
Average Price	6.300	7.045	6.625	6.316	6.426	6.724	13.861	8.466
Average OI	2685	2409	2356	2227	2234	2261	2041	2299
Average Volume	51	83	97	146	99	84	38	81
Average Volume/OI Ratio	0.02	0.03	0.04	0.07	0.04	0.04	0.02	0.04

Notes: This table reports descriptive statistics of crude oil futures and option data from January 2, 1990 to December 31, 2012. Panel A presents daily futures contract data across maturities. Panel B and C report daily crude oil option data across maturities and moneyness on each Wednesday, where moneyness is defined as the strike price K divided by the underlying futures price F.

Table 2: Average Returns of Crude Oil Futures Options

Panel A: Put option returns	1	2	3	4	5			
Moneyness	(0.5, 0.85]	(0.85, 0.95]	(0.95, 1.05]	(1.05, 1.15]	(1.15, 1.5)	(1-5)	(1-4)	(2-5)
1990-2012								
Ret	-0.881	-0.309	-0.130	-0.111	-0.056	-0.741	-0.691	-0.171
t-stat	**(-11.19)**	**(-2.04)**	(-1.48)	**(-2.11)**	(-1.47)	**(-5.02)**	**(-7.08)**	(-0.71)
10th Percentile	-1.000	-1.000	-1.000	-1.000	-0.580	-1.344	-1.851	-1.150
50th Percentile	-1.000	-1.000	-0.916	-0.236	-0.036	-0.915	-0.712	-0.709
90th Percentile	-1.000	0.648	1.826	1.035	0.534	-0.321	0.000	0.291
1990-2004								
Ret	-0.959	-0.448	-0.179	-0.149	-0.109	-0.846	-0.664	-0.484
t-stat	**(-42.24)**	**(-3.08)**	(-1.65)	**(-2.35)**	**(-2.23)**	**(-13.41)**	**(-8.32)**	**(-3.71)**
2005-2012								
Ret	-0.812	-0.071	-0.039	-0.045	0.017	-0.663	-0.712	0.251
t-stat	**(-5.54)**	(-0.22)	(-0.26)	(-0.49)	(0.28)	**(-2.62)**	**(-4.35)**	(0.47)
2005-2008								
Ret	-0.652	0.215	0.049	0.020	0.113	-0.293	-0.531	1.030
t-stat	**(-2.14)**	(0.38)	(0.23)	(0.14)	(1.00)	(-0.45)	(-1.52)	(0.89)
2009-2012								
Ret	-0.959	-0.356	-0.126	-0.102	-0.047	-0.889	-0.853	-0.277
t-stat	**(-25.99)**	(-1.09)	(-0.61)	(-0.87)	(-0.70)	**(-12.37)**	**(-7.81)**	(-0.65)

Panel B: Call option returns	1	2	3	4	5			
Moneyness	(0.5, 0.85]	(0.85, 0.95]	(0.95, 1.05]	(1.05, 1.15]	(1.15, 1.5)	(5-1)	(5-2)	(4-1)
1990-2012								
Ret	0.004	0.147	0.274	0.099	-0.774	-0.885	-0.890	-0.404
t-stat	(0.07)	**(2.38)**	**(2.41)**	(0.37)	**(-6.92)**	**(-12.59)**	**(-8.36)**	**(-2.54)**
10th Percentile	-0.759	-1.000	-1.000	-1.000	-1.000	-1.537	-2.051	-1.232
50th Percentile	0.000	0.027	-0.695	-1.000	-1.000	-0.957	-0.923	-0.835
90th Percentile	0.654	1.229	2.807	2.201	-1.000	-0.202	0.000	-0.014
1990-2004								
Ret	0.090	0.244	0.429	0.411	-0.618	-0.910	-0.838	-0.296
t-stat	(1.09)	**(2.85)**	**(2.71)**	(1.02)	**(-2.86)**	**(-7.01)**	**(-4.37)**	(-1.13)
2005-2012								
Ret	-0.078	-0.005	-0.018	-0.445	-0.935	-0.862	-0.947	-0.508
t-stat	(-1.06)	(-0.06)	(-0.13)	**(-2.64)**	**(-22.92)**	**(-13.33)**	**(-11.78)**	**(-2.72)**
2005-2008								
Ret	-0.314	0.044	0.075	-0.432	-0.929	-0.686	-0.975	-0.531
t-stat	**(-1.89)**	(0.34)	(0.37)	**(-1.87)**	**(-13.02)**	**(-4.14)**	**(-7.45)**	**(-4.52)**
2009-2012								
Ret	-0.008	-0.051	-0.110	-0.458	-0.940	-0.915	-0.922	-0.501
t-stat	(-0.10)	(-0.49)	(-0.62)	**(-1.85)**	**(-22.32)**	**(-13.65)**	**(-9.37)**	**(-2.08)**

Notes: This table reports the average returns of options across moneyness. Moneyness is defined as the strike price K divided by the futures price F. On the third Wednesday of each month, we calculate expected hold-to-maturity returns of available options, with the maturity as close to 30 calendar days as possible. Without interpolation or extrapolation, we assign available options to various bins according to their moneyness, and returns are averaged within each bin. For each moneyness interval, we show the average return, its standard t-statistics, the 10^{th}, 50^{th} and 90^{th} percentile from January 1990 to December 2012, as well as average returns for two subsample periods: January 1990 to December 2004, January 2005 to December 2012. We also report the period of 2005-2008 and 2009-2012 for comparison. Returns of call and put are calculated as $r_{t,T}^{call} = \max(F_{t,T}e^R - K, 0)/c_{t,T} - 1$ and $r_{t,T}^{put} = \max(K - F_{t,T}e^R, 0)/p_{t,T} - 1$, where $F_{t,T}$ is the price of underlying futures, $R = \log(F_T/F_{t,T})$ is the return of the underlying futures contract over the period $T - t$, K is the strike price, $c_{t,T}$ and $p_{t,T}$ are prices of European style call and put options. Standard t-statistics are reported in the parentheses and t-stats larger than 1.68 in magnitude are in boldface.

Table 3: Summary Statistics of Regression Variables

Variable	Mean	Std Dev	Percentile			P-value (ADF test)
			25th	50th	75th	
Risk-Neutral Skewness	-0.173	0.193	-0.289	-0.182	-0.078	0.001
Speculation Index	1.066	0.044	1.029	1.049	1.104	0.671
Speculation Index (Adjusted)	0.000	0.014	-0.009	-0.001	0.009	0.001
Futures Volume	123686	112482	46440	78692	147474	0.001
Futures Volume (Adjusted)	0.000	0.039	-0.013	0.000	0.013	0.001
OMT Put/Call	1.512	1.374	0.735	1.148	1.833	0.001
Option Volume	19202	17467	8382	13824	23402	0.001
Option Volume (Adjusted)	0.000	0.011	-0.005	-0.001	0.003	0.001
Bull-Bear Spread	14.252	15.236	4.700	16.200	25.300	0.001
Consumer Sentiment	86.381	13.293	76.200	88.450	94.700	0.406
Consumer Sentiment (Adjusted)	0.000	1.000	-0.577	-0.031	0.586	0.001
VIX (%)	20.952	8.913	14.400	19.265	24.865	0.012
Basis	0.016	0.200	-0.084	0.020	0.085	0.001
Storage Level	325329	25734	305639	327121	343221	0.651
Storage Level (Adjusted)	0.000	0.011	-0.007	-0.001	0.008	0.001
Volatility (%)	33.335	13.157	26.256	31.788	37.944	0.046
Historical Return (%)	0.030	1.003	-0.569	0.079	0.624	0.001

Notes: This table reports the summary statistics of variables used in the main regression analysis (14). Variables are measured as closely as possible to each Wednesday from January 1990 to December 2012. Skewness of one-month futures returns is calculated from the risk-neutral densities implied by option prices. Speculation Index measures the extent by which speculative positions exceed the necessary level to offset hedging positions, and a Hodrick-Prescott filter is applied to remove the trend. OTM Put/Call is the ratio of total open interest for OTM puts over OTM calls written on one-month futures. Futures Volume and Option Volume are the trading volume of one-month maturity crude oil futures and options, and they are adjusted for a deterministic time trend. Bull-bear Spread is the proportion of bullish investors minus bearish investors based on the survey conducted by Investors Intelligence. Consumer Sentiment is the Index of Consumer Sentiment from the University of Michigan. We take the first difference to remove the autocorrelation. The VIX is the CBOE's Volatility Index. Basis is the discounted spread between futures and spot prices. Storage Level is U.S. total stocks of crude oil, excluding SPR, based on the report from Energy Information Administration (EIA), and a Hodrick-Prescott filter is applied to remove the trend. Volatility is the 30-day option-implied volatility. Historical Return is the moving average of daily returns of holding a long position of one-month futures contracts during the previous week.

Table 4: Correlation Matrix of Regression Variables

	Spec Index	Fut Vol	Put/ Call	Opt Vol	Bull- Bear	Cons. Sent	VIX	Basis	Stor Level	Volat
Speculation Index	-									
Futures Volume	0.069									
OMT Put/Call	0.063	0.052								
Option Volume	-0.033	0.182	0.010							
Bull-Bear Spread	-0.121	0.013	0.052	0.005						
Consumer Sentiment	-0.086	-0.058	-0.104	-0.027	0.161					
VIX	0.107	-0.024	-0.045	0.003	-0.282	-0.134				
Basis	-0.035	-0.054	-0.356	0.005	-0.016	0.052	0.110			
Storage Level	-0.093	0.071	-0.008	0.055	0.013	-0.043	-0.065	0.261		
Volatility	0.031	-0.015	-0.003	0.057	-0.098	-0.023	0.648	0.175	-0.094	
Historical Return	-0.013	-0.039	0.338	-0.044	0.088	0.020	-0.073	-0.136	0.081	-0.037

Notes: This table reports the correlation matrix of variables used in the regression (14). Variables are measured as closely as possible to each Wednesday from January 1990 to December 2012. Speculation Index measures the extent by which speculative positions exceed the necessary level to offset hedging positions, and a Hodrick-Prescott filter is applied to remove the trend. OTM Put/Call is the ratio of total open interest for OTM puts over OTM calls written on one-month futures. Futures Volume and Option Volume are the trading volume of one-month maturity crude oil futures and options, and they are adjusted for a deterministic time trend. Bull-bear Spread is the proportion of bullish investors minus bearish investors based on the survey conducted by Investors Intelligence. Consumer Sentiment is the Index of Consumer Sentiment from the University of Michigan; and we take the first difference to remove the autocorrelation. The VIX is the CBOE's Volatility Index. Basis is the discounted spread between futures and spot prices. Storage Level is U.S. total stocks of crude oil, excluding SPR, based on the report from Energy Information Administration (EIA), and a Hodrick-Prescott filter is applied to remove the trend. Volatility is the 30-day option-implied volatility. Historical Return is the moving average of daily returns of holding a long position of one-month futures contracts during the previous week.

Table 5: Investor Beliefs and the Slope of the SPDs

	Left Slope (I)		Left Slope (II)		Right Slope (I)		Right Slope (II)	
	1990-2004	**2005-2012**	**1990-2004**	**2005-2012**	**1990-2004**	**2005-2012**	**1990-2004**	**2005-2012**
Speculation Index	0.006	0.007	-0.013	-0.011	0.006	-0.044	0.007	-0.057
	(0.20)	(0.31)	(-0.40)	(-0.44)	(0.21)	**(-2.07)**	(0.25)	**(-2.22)**
Futures Volume	-0.119	-0.081	-0.135	-0.090	-0.246	-0.087	-0.246	-0.105
	(-1.51)	**(-3.14)**	**(-1.71)**	**(-3.52)**	**(-3.64)**	**(-3.73)**	**(-3.10)**	**(-3.47)**
OTM Ratio	0.083	-0.043	0.056	-0.077	0.100	0.169	0.116	0.220
	(2.70)	(-0.49)	**(1.75)**	(-0.87)	(1.52)	**(2.60)**	**(2.01)**	**(2.62)**
Options Volume	-0.019	0.043	0.000	0.050	-0.035	0.035	-0.036	0.052
	(-0.29)	(1.44)	(0.00)	**(1.78)**	(-0.50)	(1.32)	(-0.44)	(1.55)
Bull-Bear Spread	-0.035	0.103	-0.076	0.0391	-0.021	-0.031	-0.003	-0.035
	(-1.04)	**(3.11)**	**(-2.64)**	(1.03)	(-0.81)	(-0.79)	(-0.11)	(-0.76)
Consumer Sentiment	0.017	-0.011	-0.015	-0.011	0.001	-0.006	0.001	-0.016
	(0.68)	(-0.49)	(-0.53)	(-0.48)	(0.05)	(-0.30)	(0.06)	(-0.68)
VIX	0.162	0.081	0.065	0.051	0.002	-0.023	0.067	-0.023
	(3.80)	(1.47)	(1.36)	(1.12)	(0.06)	(-0.58)	(1.35)	(-0.46)
Lagged Skew	0.635	0.594	0.589	0.597	0.737	0.587	0.672	0.579
	(16.96)	**(8.74)**	**(12.18)**	**(9.05)**	**(16.12)**	**(7.12)**	**(15.20)**	**(7.51)**
Obs	721	384	720	381	736	385	736	385
Adj. R^2	0.472	0.609	0.430	0.569	0.605	0.502	0.506	0.563

Notes: This table presents the results of weekly regressions that examine the dependence of the slope of the SPDs on investor beliefs. The dependent variables in the first two columns are the decreasing-region slopes of the U-shaped SPDs measured using return intervals (-0.20, -0.15) and (-0.18, -0.12); the dependent variables in the last two columns are the increasing-region slopes of the U-shaped SPDs measured using return intervals (0.10, 0.16) and (0.13, 0.18). OTM Ratio in the left (right) slope regressions are defined as the open interest ratios of OTM puts over OTM calls (ratios of OTM calls over OTM puts). For each maturity, we report the regression results of two sub-sample periods: 1990 to 2004, and 2005 to 2012. Newey-West (1987) t-statistics are reported in the parentheses. T-stats larger than 1.68 in magnitude are reported in boldface.

Table 6: Investor Beliefs and Risk-Neutral Skewness

	1990-2004				2005-2012				2005-2008			
Speculation Index	0.015		0.030	0.032	-0.026		-0.025	-0.0327	-0.0528		-0.0537	-0.0701
	(0.70)		(1.42)	(1.60)	(-1.48)		(-1.59)	**(-1.77)**	(-1.55)		(-1.65)	**(-2.25)**
Futures Volume	0.023		0.008	0.008	0.016		0.012	0.014	0.008		0.006	0.011
	(0.49)		(0.14)	(0.14)	(0.79)		(0.65)	(0.78)	(0.16)		(0.12)	(0.22)
OTM Put/Call		-0.113	-0.115	-0.114		-0.031	-0.039	-0.038		0.008	-0.002	0.010
		(-3.64)	**(-3.66)**	**(-3.61)**		(-0.63)	(-0.84)	(-0.82)		(0.16)	(-0.04)	(0.26)
Options Volume		0.063	0.063	0.065		0.033	0.030	0.030		0.041	0.042	0.036
		(2.29)	**(1.88)**	**(1.87)**		**(2.22)**	**(2.12)**	**(2.01)**		(1.59)	(1.56)	(1.10)
Bull-Bear Spread				-0.021				-0.032				-0.051
				(-0.96)				(-1.06)				(-1.06)
Consumer Sentiment				0.011				-0.003				0.004
				(0.55)				(-0.25)				(0.17)
VIX				-0.015				-0.006				0.017
				(-0.46)				(-0.23)				(0.45)
Lagged Skew	0.837	0.791	0.794	0.792	0.744	0.746	0.733	0.718	0.71	0.734	0.708	0.659
	(24.85)	**(19.01)**	**(19.19)**	**(18.48)**	**(16.27)**	**(15.96)**	**(15.88)**	**(13.94)**	**(12.16)**	**(12.20)**	**(12.11)**	**(9.68)**
Obs	740	740	740	740	392	392	392	392	203	203	203	203
Adj. R^2	0.713	0.727	0.726	0.726	0.577	0.582	0.583	0.581	0.549	0.546	0.549	0.554

Notes: This table presents the results of weekly regressions that examine the dependence of risk-neutral skewness on investor beliefs. The dependent variable is the skewness based on the risk-neutral densities implied from option prices. We report the regression results from two sub-sample periods: 1990 to 2004 and 2005 to 2012, and the period of 2005-2008, when the oil market was most volatile, for comparison. Newey-West (1987) t-statistics are reported in the parentheses. T-stats larger than 1.68 in magnitude are reported in boldface.

Table 7: Investor Beliefs and Risk-Neutral Skewness

(Robustness to Control Variables)

	Panel A: Oil Control			**Panel B**: Macro Control		
	1990-2004	2005-2012	2005-2008	1990-2004	2005-2012	2005-2008
Speculation Index	0.029	-0.034	-0.062	0.037	-0.032	-0.060
	(1.36)	**(-2.01)**	**(-2.29)**	**(1.87)**	**(-1.83)**	**(-2.21)**
Futures Volume	0.009	0.014	0.012	0.009	0.012	0.009
	(0.16)	(0.77)	(0.25)	(0.16)	(0.68)	(0.18)
OTM Put/Call	-0.108	-0.030	0.016	-0.119	-0.032	0.002
	(-3.40)	(-0.58)	(0.28)	**(-3.63)**	(-0.67)	(0.04)
Options Volume	0.046	0.031	0.038	0.071	0.033	0.053
	(1.37)	**(2.02)**	(1.23)	**(1.99)**	**(2.19)**	**(1.84)**
Bull-Bear Spread	-0.017	-0.038	-0.058	-0.013	-0.030	-0.064
	(-0.65)	(-1.12)	(-1.10)	(-0.55)	(-0.80)	(-1.06)
Consumer Sentiment	-0.003	-0.006	-0.004	0.008	-0.013	-0.010
	(-0.15)	(-0.42)	(-0.20)	(0.40)	(-1.05)	(-0.50)
VIX	-0.069	0.001	0.045	-0.018	-0.034	0.029
	(-1.76)	(0.06)	(1.30)	(-0.50)	(-0.83)	(0.61)
Lagged Skew	0.788	0.711	0.647	0.772	0.683	0.621
	(19.08)	**(13.26)**	**(9.98)**	**(15.94)**	**(12.23)**	**(9.97)**
Obs	740	392	203	740	392	203
Adj. R^2	0.732	0.582	0.555	0.729	0.586	0.558

Notes: This table reports the results of weekly regressions that examine the robustness of the dependence of risk-neutral skewness on investor beliefs. The dependent variables are the skewness based on the risk-neutral densities implied from option prices. Control variables in the crude oil market are the basis, storage level, historical returns, and option implied volatility. Macro control variables used in the right panel are the 1-month treasury-bill rate, the spread between yields on the 10-year treasury bond and the 3-month treasury bill, the difference in yields on Baa and Aaa corporate bonds, the Chicago Fed National Activity Index and the log growth in aggregate industrial production over the last 12 months. We report the regression results from two sub-sample periods: 1990 to 2004 and 2005 to 2012, and the period of 2005-2008, when the oil market was most volatile, for comparison. Newey-West (1987) t-statistics are reported in the parentheses. T-stats larger than 1.68 in magnitude are reported in boldface.

Table 8: Investor Beliefs and Risk-Neutral Skewness

(Robustness to the Alternative Measure of Skewness)

	Panel A: Oil Control			**Panel B**: Macro Control		
	1990-2004	2005-2012	2005-2008	1990-2004	2005-2012	2005-2008
Speculation Index	0.026	-0.035	-0.084	0.023	-0.036	-0.068
	(0.97)	**(-2.00)**	**(-2.63)**	(0.83)	**(-1.85)**	**(-2.29)**
Futures Volume	0.062	-0.003	0.015	0.076	-0.004	0.020
	(1.16)	(-0.15)	(0.48)	(1.39)	(-0.24)	(0.60)
OTM Put/Call	-0.114	0.033	0.002	-0.082	0.008	-0.002
	(-4.01)	(1.05)	(0.04)	**(-3.80)**	(0.25)	(-0.05)
Options Volume	-0.007	0.004	0.013	0.028	0.001	0.028
	(-0.18)	(0.28)	(0.50)	(0.68)	(0.08)	(1.11)
Bull-Bear Spread	0.017	-0.026	-0.061	0.004	0.002	-0.084
	(0.66)	(-0.86)	(-1.60)	(0.15)	(0.08)	**(-1.86)**
Consumer Sentiment	-0.021	0.011	0.005	-0.030	0.010	-0.001
	(-0.75)	(0.83)	(0.33)	(-1.01)	(0.76)	(-0.04)
VIX	-0.023	-0.024	0.005	0.082	-0.038	0.041
	(-0.79)	(-0.84)	(0.17)	**(2.46)**	(-1.12)	(1.10)
Lagged Skew	0.645	0.743	0.568	0.663	0.717	0.534
	(17.46)	**(14.94)**	**(5.54)**	**(19.65)**	**(14.23)**	**(4.80)**
Obs	740	392	203	740	392	203
Adj. R^2	0.561	0.653	0.515	0.554	0.659	0.532

Notes: This table reports the results of weekly regressions that examine the robustness of the dependence of risk-neutral skewness on investor beliefs using another measure of risk-neutral skewness. The dependent variable, risk-neutral skewness, is calculated directly from OTM option prices using the methodology of Bakshi, Kapadia, and Madan (2003). Control variables in the crude oil market are the basis, storage level, historical returns, and option implied volatility. Control variables used in the right panel are the 1-month treasury-bill rate, the spread between yields on the 10-year treasury bond and the 3-month treasury bill, the difference in yields on Baa and Aaa corporate bonds, the Chicago Fed National Activity Index and the log growth in aggregate industrial production over the last 12 months. We report the regression results from two sub-sample periods: 1990 to 2004 and 2005 to 2012, and the period of 2005-2008, when the oil market was most volatile, for comparison. Newey-West (1987) t-statistics are reported in the parentheses. T-stats larger than 1.68 in magnitude are reported in boldface.